

Two examples of queueing networks

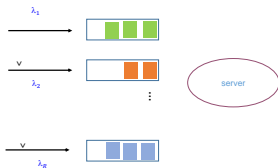
Susceptible-Infective Epidemic propagation models

Laurent Massoulié

MSR-Inria Joint Centre

October 6, 2015

Priority queue



- R customer types
- Infinite queue, single server with unit capacity
- Policy: always serve customer with highest priority (lowest class index)
Interrupt lower priority service upon higher priority arrival
Resume interrupted service where it was stopped (FIFO per class)
- Poisson λ_r arrivals in class r ; Exponential μ_r service times
Loads: $\rho_r := \lambda_r / \mu_r$

- $X_r(t)$: number of class- r customers present at time t
- A Markov jump process with only non-zero rates

$$q_{x, x+e_r} = \lambda_r, \quad q_{x, x-e_r} = \mu_r \mathbb{I}_{x_r > 0} \mathbb{I}_{x_1 = \dots = x_{r-1} = 0}$$

Proposition

Process ergodic if $\rho := \sum_r \rho_r < 1$, transient if $\rho > 1$

Assume $\mu_r \equiv \mu$ and ergodicity.

Then mean number of customers at equilibrium:

$$\mathbb{E}(X_r) = \frac{\rho_r}{(1 - \sum_{s < r} \rho_s)(1 - \sum_{s \leq r} \rho_s)}$$

Proof elements – priority queue

- Process clearly irreducible non-explosive
Foster's criterion with $V(x) := \sum_r x_r / \mu_r \Rightarrow$ ergodic if $\rho < 1$

Proof elements – priority queue

- Process clearly irreducible non-explosive
Foster's criterion with $V(x) := \sum_r x_r / \mu_r \Rightarrow$ ergodic if $\rho < 1$
- $\rho > 1$: with $X(0) = 0$, $X_r(t) = N_r(t) - D_r(t)$ (arrivals minus departures)

Work spent on class r by time t : $W_r(t) \geq \sum_{m=1}^{D_r(t)} \sigma_{r,m}$
for i.i.d. service times $\sigma_{r,m}$

Proof elements – priority queue

- Process clearly irreducible non-explosive
Foster's criterion with $V(x) := \sum_r x_r / \mu_r \Rightarrow$ ergodic if $\rho < 1$
- $\rho > 1$: with $X(0) = 0$, $X_r(t) = N_r(t) - D_r(t)$ (arrivals minus departures)
Work spent on class r by time t : $W_r(t) \geq \sum_{m=1}^{D_r(t)} \sigma_{r,m}$
for i.i.d. service times $\sigma_{r,m}$
- Law of large numbers for Poisson processes: almost surely,
 $\lim_{t \rightarrow \infty} N_r(t)/t = \lambda_r$

Proof elements – priority queue

- Process clearly irreducible non-explosive
Foster's criterion with $V(x) := \sum_r x_r / \mu_r \Rightarrow$ ergodic if $\rho < 1$
- $\rho > 1$: with $X(0) = 0$, $X_r(t) = N_r(t) - D_r(t)$ (arrivals minus departures)
Work spent on class r by time t : $W_r(t) \geq \sum_{m=1}^{D_r(t)} \sigma_{r,m}$
for i.i.d. service times $\sigma_{r,m}$
- Law of large numbers for Poisson processes: almost surely,
 $\lim_{t \rightarrow \infty} N_r(t)/t = \lambda_r$
- If for some r , $D_r(t) \leq \lambda_r t / 2$ then $X_r(t) \geq \lambda_r t / 2 + o(t)$
Else, by Law of large numbers for $\sigma_{r,m}$,
 $\forall r, W_r(t) \geq D_r(t) / \mu_r + o(t)$

Proof elements – priority queue

- Process clearly irreducible non-explosive
Foster's criterion with $V(x) := \sum_r x_r / \mu_r \Rightarrow$ ergodic if $\rho < 1$
- $\rho > 1$: with $X(0) = 0$, $X_r(t) = N_r(t) - D_r(t)$ (arrivals minus departures)

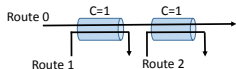
Work spent on class r by time t : $W_r(t) \geq \sum_{m=1}^{D_r(t)} \sigma_{r,m}$
for i.i.d. service times $\sigma_{r,m}$

- Law of large numbers for Poisson processes: almost surely,
 $\lim_{t \rightarrow \infty} N_r(t)/t = \lambda_r$
- If for some r , $D_r(t) \leq \lambda_r t / 2$ then $X_r(t) \geq \lambda_r t / 2 + o(t)$
Else, by Law of large numbers for $\sigma_{r,m}$,
 $\forall r, W_r(t) \geq D_r(t) / \mu_r + o(t)$
- Since $\sum_r W_r(t) \leq t$, implies

$$\sum_r X_r(t) / \mu_r \geq \rho t - t + o(t)$$

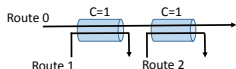
In both cases $\max_r X_r(t) \rightarrow \infty$ almost surely

Exactly solvable bandwidth sharing network



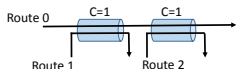
- Linear network with L unit capacity links, $L + 1$ classes. Class 0 uses all links, class r uses only link r , $r \geq 1$
- Poisson ν_r arrivals, Exponential μ_r service times, loads $\rho_r = \nu_r / \mu_r$
- x_r : number of ongoing type r transfers

Exactly solvable bandwidth sharing network



- Linear network with L unit capacity links, $L + 1$ classes. Class 0 uses all links, class r uses only link r , $r \geq 1$
- Poisson ν_r arrivals, Exponential μ_r service times, loads $\rho_r = \nu_r / \mu_r$
- x_r : number of ongoing type r transfers
- Proportionally fair allocations: service rate to class 0 : $\Lambda_0 = x_0 / [x_0 + y]$ where $y = \sum_{r=1}^L x_r$; service to class $r \geq 1$: $\Lambda_r = y / (y + x_0) \mathbb{I}_{x_r > 0}$

Exactly solvable bandwidth sharing network



- Linear network with L unit capacity links, $L + 1$ classes. Class 0 uses all links, class r uses only link r , $r \geq 1$
- Poisson ν_r arrivals, Exponential μ_r service times, loads $\rho_r = \nu_r / \mu_r$
- x_r : number of ongoing type r transfers
- Proportionally fair allocations: service rate to class 0 : $\Lambda_0 = x_0 / [x_0 + y]$ where $y = \sum_{r=1}^L x_r$; service to class $r \geq 1$: $\Lambda_r = y / (y + x_0) \mathbb{I}_{x_r > 0}$

See Problem 2, PC 3: proportionally fair shares at “macroscopic” (transmission) level result from simple processor sharing at “microscopic” (data packet) level

Exactly solvable bandwidth sharing network

- Markov jump process with non-zero rates

$$q_{x,x+e_r} = \nu_r, \quad q_{x,x-e_r} = \mu_r \Lambda_r$$

Exactly solvable bandwidth sharing network

- Markov jump process with non-zero rates

$$q_{x,x+e_r} = \nu_r, \quad q_{x,x-e_r} = \mu_r \Lambda_r$$

- reversible process for measure $\pi(x) := \binom{y+x_0}{y} \prod_{r \geq 0} \rho_r^{x_r}$

Exactly solvable bandwidth sharing network

- Markov jump process with non-zero rates

$$q_{x,x+e_r} = \nu_r, \quad q_{x,x-e_r} = \mu_r \wedge_r$$

- reversible process for measure $\pi(x) := \binom{y+x_0}{y} \prod_{r \geq 0} \rho_r^{x_r}$

Negative binomial formula: $\sum_{x_0 \geq 0} \binom{y+x_0}{y} \rho_0^{x_0} = (1 - \rho_0)^{-y-1}$

Exactly solvable bandwidth sharing network

- Markov jump process with non-zero rates

$$q_{x,x+e_r} = \nu_r, \quad q_{x,x-e_r} = \mu_r \Lambda_r$$

- reversible process for measure $\pi(x) := \binom{y+x_0}{y} \prod_{r \geq 0} \rho_r^{x_r}$

Negative binomial formula: $\sum_{x_0 \geq 0} \binom{y+x_0}{y} \rho_0^{x_0} = (1 - \rho_0)^{-y-1}$

$\Rightarrow \pi$ summable (hence process ergodic) if and only if $\rho_0 + \rho_r < 1, r = 1, \dots, L$ yielding stationary distribution

$$\pi(x) = \binom{y+x_0}{x_0} (1 - \rho_0)^{-L+1} \prod_{r=1}^L (1 - \rho_0 - \rho_r) \prod_{r=0}^L \rho_r^{x_r}$$

Exactly solvable bandwidth sharing network

- Markov jump process with non-zero rates

$$q_{x,x+e_r} = \nu_r, \quad q_{x,x-e_r} = \mu_r \Lambda_r$$

- reversible process for measure $\pi(x) := \binom{y+x_0}{y} \prod_{r \geq 0} \rho_r^{x_r}$

Negative binomial formula: $\sum_{x_0 \geq 0} \binom{y+x_0}{y} \rho_0^{x_0} = (1 - \rho_0)^{-y-1}$

$\Rightarrow \pi$ summable (hence process ergodic) if and only if $\rho_0 + \rho_r < 1, r = 1, \dots, L$ yielding stationary distribution

$$\pi(x) = \binom{y+x_0}{x_0} (1 - \rho_0)^{-L+1} \prod_{r=1}^L (1 - \rho_0 - \rho_r) \prod_{r=0}^L \rho_r^{x_r}$$

- Generating function (z-transform):

$$\mathbb{E} \prod_{r=0}^L z_r^{X_r} = \frac{(1 - \rho_0 z_0)^{L-1}}{(1 - \rho_0)^{L-1}} \prod_{r=1}^L \frac{1 - \rho_0 - \rho_r}{1 - \rho_0 z_0 - \rho_r z_r}$$

Exactly solvable bandwidth sharing network

- Markov jump process with non-zero rates

$$q_{x,x+e_r} = \nu_r, \quad q_{x,x-e_r} = \mu_r \Lambda_r$$

- reversible process for measure $\pi(x) := \binom{y+x_0}{y} \prod_{r \geq 0} \rho_r^{x_r}$

Negative binomial formula: $\sum_{x_0 \geq 0} \binom{y+x_0}{y} \rho_0^{x_0} = (1 - \rho_0)^{-y-1}$

$\Rightarrow \pi$ summable (hence process ergodic) if and only if $\rho_0 + \rho_r < 1, r = 1, \dots, L$ yielding stationary distribution

$$\pi(x) = \binom{y+x_0}{x_0} (1 - \rho_0)^{-L+1} \prod_{r=1}^L (1 - \rho_0 - \rho_r) \prod_{r=0}^L \rho_r^{x_r}$$

- Generating function (z-transform):

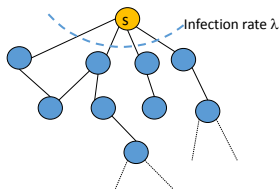
$$\mathbb{E} \prod_{r=0}^L z_r^{X_r} = \frac{(1 - \rho_0 z_0)^{L-1}}{(1 - \rho_0)^{L-1}} \prod_{r=1}^L \frac{1 - \rho_0 - \rho_r}{1 - \rho_0 z_0 - \rho_r z_r}$$

- Yields explicit formulas for per class generating functions, e.g. X_r Geometric ($\rho_r / (1 - \rho_0)$) for $r \geq 1$, and

$$\mathbb{E}(X_r) = \frac{\rho_r}{1 - \rho_0 - \rho_r}, \quad r \geq 1,$$

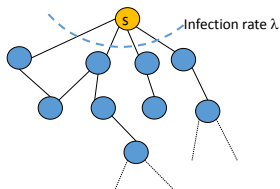
$$\mathbb{E}(X_0) = \frac{\rho_0}{1 - \rho_0} \left[1 + \sum_{r=1}^L \frac{\rho_r}{1 - \rho_0 - \rho_r} \right]$$

Susceptible-Infective epidemic propagation



- Graph $G = (V, E)$ with n nodes ($V = [n]$)
- Infected node makes infection attempts at instants of Poisson λ process, towards graph neighbor chosen uniformly at random
- Keeps attempting forever

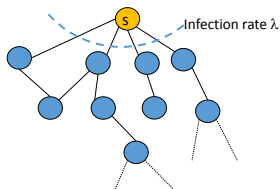
Susceptible-Infective epidemic propagation



- Graph $G = (V, E)$ with n nodes ($V = [n]$)
- Infected node makes infection attempts at instants of Poisson λ process, towards graph neighbor chosen uniformly at random
- Keeps attempting forever

⇒ Average time to total infection? Fluctuations around average? Impact of graph topology?

Susceptible-Infective epidemic propagation

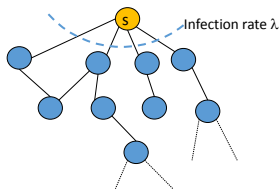


- Graph $G = (V, E)$ with n nodes ($V = [n]$)
- Infected node makes infection attempts at instants of Poisson λ process, towards graph neighbor chosen uniformly at random
- Keeps attempting forever

⇒ Average time to total infection? Fluctuations around average? Impact of graph topology?

- Variant: each node = origin of its own specific epidemics; each propagation: forwards all epidemics currently held

Susceptible-Infective epidemic propagation



- Graph $G = (V, E)$ with n nodes ($V = [n]$)
- Infected node makes infection attempts at instants of Poisson λ process, towards graph neighbor chosen uniformly at random
- Keeps attempting forever

\Rightarrow Average time to total infection? Fluctuations around average? Impact of graph topology?

- Variant: each node = origin of its own specific epidemics; each propagation: forwards all epidemics currently held
 \Rightarrow Time till everyone heard from everyone else (“all-to-all broadcast”)?

Susceptible-Infective epidemic propagation

- Running assumption: **complete graph**

Susceptible-Infective epidemic propagation

- Running assumption: **complete graph**
- System description: X_t = number of infected nodes at time t

Susceptible-Infective epidemic propagation

- Running assumption: **complete graph**
- System description: X_t = number of infected nodes at time t
- Rate at which new attempts made when in state x :
superposition of x Poisson λ processes
Success probability of infection attempt: $(n - x)/(n - 1)$

Susceptible-Infective epidemic propagation

- Running assumption: **complete graph**
- System description: X_t = number of infected nodes at time t
- Rate at which new attempts made when in state x :
superposition of x Poisson λ processes
Success probability of infection attempt: $(n - x)/(n - 1)$

⇒ next infection time: first time of Poisson $x\lambda$ process,
thinned with probability $(n - x)/(n - 1)$ of retaining points:
Exponential $\lambda x(n - x)/(n - 1)$ random variable

Susceptible-Infective epidemic propagation

- Running assumption: **complete graph**
- System description: X_t = number of infected nodes at time t
- Rate at which new attempts made when in state x :
superposition of x Poisson λ processes
Success probability of infection attempt: $(n - x)/(n - 1)$

\Rightarrow next infection time: first time of Poisson $x\lambda$ process,
thinned with probability $(n - x)/(n - 1)$ of retaining points:
Exponential $\lambda x(n - x)/(n - 1)$ random variable

$\Rightarrow X_t$ a Markov jump process with non-zero jump rate
 $q_{x,x+1} = \lambda x(n - x)/(n - 1)$

Time to total infection

Let E_x : i.i.d. Exponential(1) random variables, T_n : time to total outbreak

Then $T_n = \sum_{x=1}^{n-1} \frac{1}{q_x} E_x$, with $q_x = \lambda x(n-x)/(n-1)$

Time to total infection

Let E_x : i.i.d. Exponential(1) random variables, T_n : time to total outbreak

Then $T_n = \sum_{x=1}^{n-1} \frac{1}{q_x} E_x$, with $q_x = \lambda x(n-x)/(n-1)$

$$\begin{aligned}\mathbb{E}(T_n) &= \sum_{x=1}^{n-1} \frac{1}{q_x} = \frac{n-1}{n} \frac{1}{\lambda} \sum_{x=1}^{n-1} \left(\frac{1}{x} + \frac{1}{n-x} \right) \\ &= \frac{n-1}{n} \frac{2}{\lambda} H(n-1) \\ &= \frac{2}{\lambda} [\ln(n) + \gamma + o(1)]\end{aligned}$$

where $H(k)$: k -th Harmonic number, and $\gamma \approx 0.577$: Euler's constant

Time to total infection

Let E_x : i.i.d. Exponential(1) random variables, T_n : time to total outbreak

Then $T_n = \sum_{x=1}^{n-1} \frac{1}{q_x} E_x$, with $q_x = \lambda x(n-x)/(n-1)$

$$\begin{aligned}\mathbb{E}(T_n) &= \sum_{x=1}^{n-1} \frac{1}{q_x} = \frac{n-1}{n} \frac{1}{\lambda} \sum_{x=1}^{n-1} \left(\frac{1}{x} + \frac{1}{n-x} \right) \\ &= \frac{n-1}{n} \frac{2}{\lambda} H(n-1) \\ &= \frac{2}{\lambda} [\ln(n) + \gamma + o(1)]\end{aligned}$$

where $H(k)$: k -th Harmonic number, and $\gamma \approx 0.577$: Euler's constant

Similarly, for $0 < a < b < 1$: $\mathbb{E}(T_{bn} - T_{an}) \rightarrow \frac{1}{\lambda} \ln \left(\frac{b}{1-b} \frac{1-a}{a} \right)$

Time to total infection

Let E_x : i.i.d. Exponential(1) random variables, T_n : time to total outbreak

Then $T_n = \sum_{x=1}^{n-1} \frac{1}{q_x} E_x$, with $q_x = \lambda x(n-x)/(n-1)$

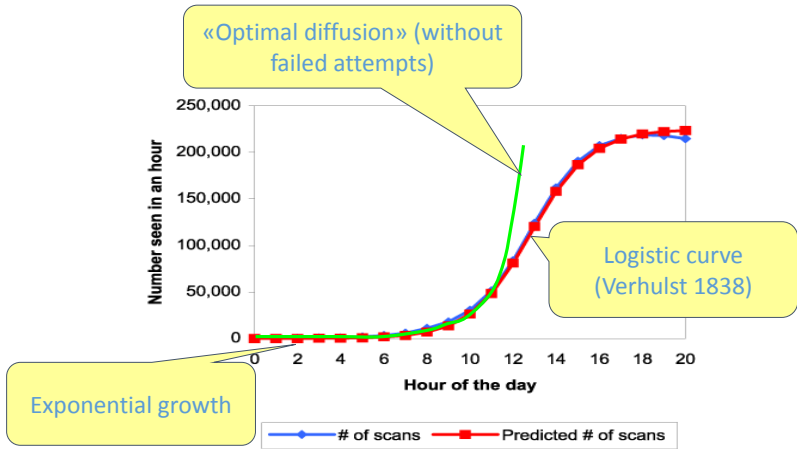
$$\begin{aligned}\mathbb{E}(T_n) &= \sum_{x=1}^{n-1} \frac{1}{q_x} = \frac{n-1}{n} \frac{1}{\lambda} \sum_{x=1}^{n-1} \left(\frac{1}{x} + \frac{1}{n-x} \right) \\ &= \frac{n-1}{n} \frac{2}{\lambda} H(n-1) \\ &= \frac{2}{\lambda} [\ln(n) + \gamma + o(1)]\end{aligned}$$

where $H(k)$: k -th Harmonic number, and $\gamma \approx 0.577$: Euler's constant

Similarly, for $0 < a < b < 1$: $\mathbb{E}(T_{bn} - T_{an}) \rightarrow \frac{1}{\lambda} \ln \left(\frac{b}{1-b} \frac{1-a}{a} \right)$

Heuristic inversion: starting from $X_0 = an$, $X_t \approx n \frac{ae^{\lambda t}}{1-a+ae^{\lambda t}}$

\Rightarrow The celebrated **logistic function**, or S-curve



Time to total infection order-optimal
(logarithmic in number of targets) despite random target selection

Controlling fluctuations

- **Markov's inequality:** random variable $X \geq 0$,
 $a > 0 \Rightarrow \mathbb{P}(X \geq a) \leq \mathbb{E}(X)/a$

Controlling fluctuations

- **Markov's inequality:** random variable $X \geq 0$,
 $a > 0 \Rightarrow \mathbb{P}(X \geq a) \leq \mathbb{E}(X)/a$
- **Bienaymé-Tchebitchev's inequality:** random variable
 $X \in \mathbb{R}$: $\mathbb{P}(|X - \mathbb{E}(X)| \geq a) \leq \text{Var}(X)/a^2$

Controlling fluctuations

- **Markov's inequality:** random variable $X \geq 0$,
 $a > 0 \Rightarrow \mathbb{P}(X \geq a) \leq \mathbb{E}(X)/a$
- **Bienaymé-Tchebitchev's inequality:** random variable
 $X \in \mathbb{R}$: $\mathbb{P}(|X - \mathbb{E}(X)| \geq a) \leq \text{Var}(X)/a^2$
- **Exponential version:** for $\theta > 0$, $\mathbb{P}(X \geq t) \leq \mathbb{E}(e^{\theta X})e^{-\theta t}$ i.e.
finite exponential moments yield exponentially decaying
control of tail probabilities

Controlling fluctuations

- **Markov's inequality:** random variable $X \geq 0$,
 $a > 0 \Rightarrow \mathbb{P}(X \geq a) \leq \mathbb{E}(X)/a$
- **Bienaymé-Tchebitchev's inequality:** random variable
 $X \in \mathbb{R}$: $\mathbb{P}(|X - \mathbb{E}(X)| \geq a) \leq \text{Var}(X)/a^2$
- **Exponential version:** for $\theta > 0$, $\mathbb{P}(X \geq t) \leq \mathbb{E}(e^{\theta X})e^{-\theta t}$ i.e.
finite exponential moments yield exponentially decaying
control of tail probabilities

Variable $S_n := \lambda(T_n - \mathbb{E}(T_n))$ satisfies for all $\theta \in [0, 1/2]$
 $\mathbb{E}(\exp(\theta S_n)) \leq \exp(4\pi^2\theta^2/3) =: C_\theta < +\infty$
hence $\mathbb{P}(\lambda(T_n - \mathbb{E}(T_n)) \geq t) \leq C_\theta e^{-\theta t}$ (fluctuations small
compared to mean)

Controlling fluctuations

- **Markov's inequality:** random variable $X \geq 0$,
 $a > 0 \Rightarrow \mathbb{P}(X \geq a) \leq \mathbb{E}(X)/a$
- **Bienaymé-Tchebitchev's inequality:** random variable
 $X \in \mathbb{R}$: $\mathbb{P}(|X - \mathbb{E}(X)| \geq a) \leq \text{Var}(X)/a^2$
- **Exponential version:** for $\theta > 0$, $\mathbb{P}(X \geq t) \leq \mathbb{E}(e^{\theta X})e^{-\theta t}$ i.e.
finite exponential moments yield exponentially decaying
control of tail probabilities

Variable $S_n := \lambda(T_n - \mathbb{E}(T_n))$ satisfies for all $\theta \in [0, 1/2]$
 $\mathbb{E}(\exp(\theta S_n)) \leq \exp(4\pi^2\theta^2/3) =: C_\theta < +\infty$
hence $\mathbb{P}(\lambda(T_n - \mathbb{E}(T_n)) \geq t) \leq C_\theta e^{-\theta t}$ (fluctuations small
compared to mean)

Proof: For $r_x = x(n-x)/(n-1) = q_x/\lambda$,

$$\mathbb{E}e^{\theta S_n} = \prod_{x=1}^{n-1} \frac{r_x}{r_x - \theta} e^{-\theta/r_x}$$

For $u \in (0, 1/2]$, $\frac{e^{-u}}{1-u} \leq 1 + 2u^2$, hence:

$$\mathbb{E}e^{\theta S_n} \leq \prod_{x=1}^{n-1} [1 + 2(\theta/r_x)^2] \leq e^{\sum_{x=1}^{n-1} 2(\theta/r_x)^2} \leq e^{8\theta^2 \sum_{x \geq 1} x^{-2}}$$

Application: All-to-all scenario (one epidemic per user)

Lemma

Let random variables S^1, \dots, S^n be such that for some $a, b > 0$:

$$\forall t > 0, \forall i \in [n], \mathbb{P}(S^i \geq t) \leq ae^{-bt}$$

$$\text{Then } \mathbb{E}(\sup_i S^i) \leq \mathbb{E}((\sup_i S^i)^+) \leq \frac{\ln(an)+1}{b}$$

Application: All-to-all scenario (one epidemic per user)

Lemma

Let random variables S^1, \dots, S^n be such that for some $a, b > 0$:

$$\forall t > 0, \forall i \in [n], \mathbb{P}(S^i \geq t) \leq ae^{-bt}$$

$$\text{Then } \mathbb{E}(\sup_i S^i) \leq \mathbb{E}((\sup_i S^i)^+) \leq \frac{\ln(an)+1}{b}$$

Proof: Write $\mathbb{E}((\sup_i S^i)^+) = \int_0^\infty \mathbb{P}(\sup_i S^i \geq t) dt$

Then upper-bound $\mathbb{P}(\sup_i S^i \geq t)$ by nae^{-bt} for $t \geq \ln(an)/b$ (**union bound**) and by 1 otherwise.

Application: All-to-all scenario (one epidemic per user)

Lemma

Let random variables S^1, \dots, S^n be such that for some $a, b > 0$:

$$\forall t > 0, \forall i \in [n], \mathbb{P}(S^i \geq t) \leq ae^{-bt}$$

$$\text{Then } \mathbb{E}(\sup_i S^i) \leq \mathbb{E}((\sup_i S^i)^+) \leq \frac{\ln(an)+1}{b}$$

Proof: Write $\mathbb{E}((\sup_i S^i)^+) = \int_0^\infty \mathbb{P}(\sup_i S^i \geq t) dt$

Then upper-bound $\mathbb{P}(\sup_i S^i \geq t)$ by nae^{-bt} for $t \geq \ln(an)/b$ (**union bound**) and by 1 otherwise.

Corollary

All-to-all propagation time T satisfies for all $\theta \in (0, 1/2]$

$$\mathbb{E}T \leq \frac{1}{\lambda} \left[2(\ln(n) + \gamma) + o(1) + \frac{\ln(C_\theta n)+1}{\theta} \right] = O(\ln(n)),$$

same order still

Application: All-to-all scenario (one epidemic per user)

Lemma

Let random variables S^1, \dots, S^n be such that for some $a, b > 0$:

$$\forall t > 0, \forall i \in [n], \mathbb{P}(S^i \geq t) \leq ae^{-bt}$$

$$\text{Then } \mathbb{E}(\sup_i S^i) \leq \mathbb{E}((\sup_i S^i)^+) \leq \frac{\ln(an)+1}{b}$$

Proof: Write $\mathbb{E}((\sup_i S^i)^+) = \int_0^\infty \mathbb{P}(\sup_i S^i \geq t) dt$

Then upper-bound $\mathbb{P}(\sup_i S^i \geq t)$ by nae^{-bt} for $t \geq \ln(an)/b$ (**union bound**) and by 1 otherwise.

Corollary

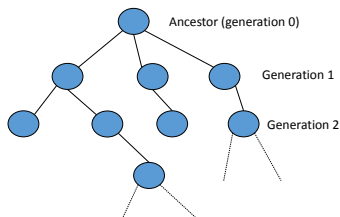
All-to-all propagation time T satisfies for all $\theta \in (0, 1/2]$

$$\mathbb{E}T \leq \frac{1}{\lambda} \left[2(\ln(n) + \gamma) + o(1) + \frac{\ln(C_\theta n)+1}{\theta} \right] = O(\ln(n)),$$

same order still

Indeed: $T =$ supremum of n propagation times corresponding each to single epidemic propagation

Towards Susceptible-Infective-Removed (SIR) epidemics: Galton-Watson branching process (1873)



Offspring distribution $\{p_k\}_{k \in \mathbb{N}}$

Z_k number of individuals per generation:

$$Z_0 = 1, Z_k = \sum_{m=1}^{Z_{k-1}} X_{m,k} \text{ where } \{X_{m,k}\}_{m,k \geq 0}: \text{i.i.d.}, \sim \{p_k\}_{k \in \mathbb{N}}$$

Quantities of interest: probability of extinction; in case of extinction, total population size

Theorem

Extinction probability p_{ext} : smallest root in $[0, 1]$ of $z = \phi(z)$

where $\phi(z) = \mathbb{E}(z^X) = \sum_{k \geq 0} p_k z^k$

If $\mu := \mathbb{E}(X) < 1$ then $p_{\text{ext}} = 1$

If $\mu = 1$ and $p_0 > 0$ then $p_{\text{ext}} = 1$

If $\mu > 1$ then $p_{\text{ext}} < 1$

Theorem

Extinction probability p_{ext} : smallest root in $[0, 1]$ of $z = \phi(z)$

where $\phi(z) = \mathbb{E}(z^X) = \sum_{k \geq 0} p_k z^k$

If $\mu := \mathbb{E}(X) < 1$ then $p_{\text{ext}} = 1$

If $\mu = 1$ and $p_0 > 0$ then $p_{\text{ext}} = 1$

If $\mu > 1$ then $p_{\text{ext}} < 1$

Proof: $\{Z_k = 0\} \nearrow \{\text{Extinction}\}$; $\mathbb{P}(Z_k = 0) = \phi_k(0)$ where
 $\phi_k(z) = \mathbb{E}(z^{Z_k})$

By induction $\phi_k(z) = \phi \circ \phi_{k-1}(z)$ hence

$\mathbb{P}(Z_k = 0) = \phi(\mathbb{P}(Z_{k-1} = 0))$

\Rightarrow by monotonicity of ϕ and $\mathbb{P}(Z_0 = 0) = 0$, sequence increases to
(necessarily smallest) fixed point.

Theorem

Extinction probability p_{ext} : smallest root in $[0, 1]$ of $z = \phi(z)$

where $\phi(z) = \mathbb{E}(z^X) = \sum_{k \geq 0} p_k z^k$

If $\mu := \mathbb{E}(X) < 1$ then $p_{\text{ext}} = 1$

If $\mu = 1$ and $p_0 > 0$ then $p_{\text{ext}} = 1$

If $\mu > 1$ then $p_{\text{ext}} < 1$

Proof: $\{Z_k = 0\} \nearrow \{\text{Extinction}\}$; $\mathbb{P}(Z_k = 0) = \phi_k(0)$ where
 $\phi_k(z) = \mathbb{E}(z^{Z_k})$

By induction $\phi_k(z) = \phi \circ \phi_{k-1}(z)$ hence

$\mathbb{P}(Z_k = 0) = \phi(\mathbb{P}(Z_{k-1} = 0))$

\Rightarrow by monotonicity of ϕ and $\mathbb{P}(Z_0 = 0) = 0$, sequence increases to
(necessarily smallest) fixed point.

μ : slope of ϕ at 1^- . By convexity of ϕ , only fixed point: 1 if $\mu < 1$

By continuity of ϕ , \exists fixed point < 1 if $\mu > 1$

For $\mu = 1$, if $p_0 > 0$ then ϕ strictly convex hence only fixed point:

1; if $p_0 = 0$ then $p_{\text{ext}} = 0$

Theorem

Extinction probability p_{ext} : smallest root in $[0, 1]$ of $z = \phi(z)$

where $\phi(z) = \mathbb{E}(z^X) = \sum_{k \geq 0} p_k z^k$

If $\mu := \mathbb{E}(X) < 1$ then $p_{\text{ext}} = 1$

If $\mu = 1$ and $p_0 > 0$ then $p_{\text{ext}} = 1$

If $\mu > 1$ then $p_{\text{ext}} < 1$

Proof: $\{Z_k = 0\} \nearrow \{\text{Extinction}\}$; $\mathbb{P}(Z_k = 0) = \phi_k(0)$ where
 $\phi_k(z) = \mathbb{E}(z^{Z_k})$

By induction $\phi_k(z) = \phi \circ \phi_{k-1}(z)$ hence

$\mathbb{P}(Z_k = 0) = \phi(\mathbb{P}(Z_{k-1} = 0))$

\Rightarrow by monotonicity of ϕ and $\mathbb{P}(Z_0 = 0) = 0$, sequence increases to
(necessarily smallest) fixed point.

μ : slope of ϕ at 1^- . By convexity of ϕ , only fixed point: 1 if $\mu < 1$

By continuity of ϕ , \exists fixed point < 1 if $\mu > 1$

For $\mu = 1$, if $p_0 > 0$ then ϕ strictly convex hence only fixed point:

1; if $p_0 = 0$ then $p_{\text{ext}} = 0$

Fundamental example of **phase transition**

Special case $X \sim \text{Poisson}(\mu)$: $p_{\text{ext}} = e^{-\mu(1-p_{\text{ext}})}$

Random walk exploration of Galton-Watson tree

Sequentially pick *active* node (whose children have not yet been sampled)

De-activate it and add its children to active set

Stop when active set empty (tree exploration complete)

Random walk exploration of Galton-Watson tree

Sequentially pick *active* node (whose children have not yet been sampled)

De-activate it and add its children to active set

Stop when active set empty (tree exploration complete)

- Dynamics of A_t , number of active nodes at step t :
Random walk $A_t = A_{t-1} - 1 + X_t$ where X_t independent of past exploration $\{A_s, X_s, s < t\}$ and distributed according to $\{p_k\}_{k \geq 0}$
- Time T at which exploration stops, i.e. $A_T = 0$ gives size of tree. Indeed $A_t = 1 - t + X_1 + \dots + X_t$ and $A_T = 0$ yield $T = 1 + X_1 + \dots + X_T$.
- Random walk can be pursued after time T

Random walk exploration of Galton-Watson tree

Sequentially pick *active* node (whose children have not yet been sampled)

De-activate it and add its children to active set

Stop when active set empty (tree exploration complete)

- Dynamics of A_t , number of active nodes at step t :
Random walk $A_t = A_{t-1} - 1 + X_t$ where X_t independent of past exploration $\{A_s, X_s, s < t\}$ and distributed according to $\{p_k\}_{k \geq 0}$
- Time T at which exploration stops, i.e. $A_T = 0$ gives size of tree. Indeed $A_t = 1 - t + X_1 + \dots + X_t$ and $A_T = 0$ yield $T = 1 + X_1 + \dots + X_T$.
- Random walk can be pursued after time T

\Rightarrow Bound on population size: for continued RW $\{A_t\}_{t \geq 0}$,

$$\mathbb{P}(T > t) = \mathbb{P}(A_0, \dots, A_t > 0) \leq \mathbb{P}(A_t > 0) = \mathbb{P}(\sum_{s=1}^t (X_s - 1) \geq 0)$$

Theorem

For i.i.d. X_s , $\mathbb{P}(\sum_{s=1}^t X_s \geq at) \leq e^{-th(a)}$ where
 $h(a) := \sup_{\theta > 0} [\theta a - \ln(\mathbb{E}(e^{\theta X_1}))]$

Theorem

For i.i.d. X_s , $\mathbb{P}(\sum_{s=1}^t X_s \geq at) \leq e^{-th(a)}$ where
 $h(a) := \sup_{\theta > 0} [\theta a - \ln(\mathbb{E}(e^{\theta X_1}))]$

Non-trivial exponential bound when $a > \mathbb{E}(X_1)$ and
 $\exists \epsilon > 0 : \mathbb{E}e^{\epsilon X_1} < +\infty$

Theorem

For i.i.d. X_s , $\mathbb{P}(\sum_{s=1}^t X_s \geq at) \leq e^{-th(a)}$ where
 $h(a) := \sup_{\theta > 0} [\theta a - \ln(\mathbb{E}(e^{\theta X_1}))]$

Non-trivial exponential bound when $a > \mathbb{E}(X_1)$ and

$$\exists \epsilon > 0 : \mathbb{E}e^{\epsilon X_1} < +\infty$$

Application to Galton-Watson process:

$\mathbb{P}(T > t) \leq e^{-th(1)}$ exponentially decaying if $\mathbb{E}(X_1) < 1$ and X_1 admits finite exponential moments.

Theorem

For i.i.d. X_s , $\mathbb{P}(\sum_{s=1}^t X_s \geq at) \leq e^{-th(a)}$ where
 $h(a) := \sup_{\theta > 0} [\theta a - \ln(\mathbb{E}(e^{\theta X_1}))]$

Non-trivial exponential bound when $a > \mathbb{E}(X_1)$ and

$$\exists \epsilon > 0 : \mathbb{E}e^{\epsilon X_1} < +\infty$$

Application to Galton-Watson process:

$\mathbb{P}(T > t) \leq e^{-th(1)}$ exponentially decaying if $\mathbb{E}(X_1) < 1$ and X_1 admits finite exponential moments.

Case of Poisson random variables, parameter $\mu > 0$, $a > \mu$:

$$h_\mu(a) = \sup_{\theta > 0} [\theta a - \mu(e^\theta - 1)]$$

Gives $\theta = \ln(a/\mu)$, $h_\mu(a) = \mu h_1(a/\mu)$

with $h_1(x) = x \ln(x) - x + 1$

Takeaway messages

- Epidemic spread in logarithmic time for single propagation and for all-to-all propagation
- Same order as if infection attempts were optimized
- Motivates “epidemic algorithms” for information dissemination
- Exponential versions of Markov’s inequality (in particular Chernoff’s inequality): powerful tool, will be used in analysis of SIR epidemics and random graphs