# **Diameter in weighted random graphs**

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# WEIGHTED DIAMETER

Graph G = (V, E):

- Distance dist $(a, b) = \min_{\pi \in \Pi(a, b)} |\pi|$ , the number of edges in E in the shortest path connecting a and b.
- Diameter of G defined by:

 $diam(G) = \max\{dist(a, b), a, b \in V, dist(a, b) < \infty\}.$ 

- Weight associated to each edge  $e \in E$ :  $w_e$ .
- Weighted distance dist $_w(a, b) = \min_{\pi \in \Pi(a, b)} \sum_{e \in \pi} w_e$ .
- Weighted diameter of G defined by:

$$\mathsf{diam}_w(G) = \max\{\mathsf{dist}_w(a,b), \ a,b \in V, \ \mathsf{dist}_w(a,b) < \infty\}$$
 .

# **CONFIGURATION MODEL**

For  $n \in \mathbb{N}$ , let  $(d_i)_1^n$  be a sequence of non-negative integers such that  $\sum_{i=1}^n d_i$  is even. We define  $G^*(n, (d_i)_1^n)$  a random multigraph with given degree sequence  $(d_i)_1^n$ : to each node i we associate  $d_i$  labeled half-edges. All half-edges need to be paired to construct the graph, this is done by randomly matching them. When a half-edge of i is paired with a half-edge of j, we interpret this as an edge between i and j.

The graph  $G^*(n, (d_i)_1^n)$  obtained following this procedure may not be simple. Conditional on the multigraph  $G^*(n, (d_i)_1^n)$  being a simple graph, we obtain a uniformly distributed random graph with the given degree sequence, which we denote by  $G(n, (d_i)_1^n)$ ,

#### **ASSUMPTIONS ON THE DEGREE SEQUENCE**

- (i) There exists a distribution  $p = \{p_k\}_{k=0}^{\infty}$  such that  $|\{i, d_i = k\}|/n \to p_k$  for every  $k \ge 0$  as  $n \to \infty$ ;
- (ii)  $\lambda := \sum_{k \ge 0} k p_k \in (0,\infty);$ (iii)  $\sum_{i=1}^n d_i^2 = O(n),$
- (iv) for some  $\tau > 0$ ,  $\Delta_n := \max_{i \in V} d_i = O(n^{1/2-\tau})$ .

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Let  $\hat{D}_n$  be the degree of a random vertex of  $G^*(n, (d_i)_1^n)$  and D a random variable with distribution  $p_k$ , then (i) is equivalent to  $\hat{D}_n \stackrel{d}{\to} D$ .

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(iii) ensures that  $\liminf \mathbb{P}(G^*(n, (d_i)_1^n) \text{ is simple}) > 0.$ 





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**EXPLORATION PROCESS** 



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#### **BRANCHING PROCESS APPROXIMATION**

The first individual has offspring distribution  $\{p_k\}$ .

The other individuals have offspring distribution  $\{q_k\}$ .

Let  $\{q_k\}_{k=0}^{\infty}$  the size-biased probability mass function corresponding to  $\{p_k\}$ , by

$$q_k = rac{(k+1)p_{k+1}}{\lambda}$$
, and,  $u = \sum_{k=0}^{\infty} kq_k \in (0,\infty).$ 

The mean of the size of generation k is  $\lambda \nu^{k-1}$ .

The condition  $\nu > 1$  is equivalent to the existence of a giant component.

## **TYPICAL GRAPH DISTANCE**

**Theorem 1.** For *a* and *b* chosen uniformly at random in the giant component of  $G(n, (d_i)_1^n)$ , we have

$$\frac{\operatorname{dist}(a,b)}{\log n} \xrightarrow{p} \frac{1}{\log \nu}.$$

Van der Hofstad, Hooghiemstra, Van Mieghem 2005

# A SIMPLE HEURISTIC

Let  $Z_k^{(1)}$  be the number of free half-edges in the ball  $B(a,k) = \{i, \operatorname{dist}(1,i) \le k\}$ .  $Z_0^{(1)}$  is the degree of node 1.

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A free half-edge of  $Z_k^{(1)}$  is attached to a free half-edge of  $Z_k^{(2)}$  with positive probability if  $Z_k^{(1)} Z_k^{(2)}$  is of order the total number of free half-edges left after k exploration steps.

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The typical distance between 1 and 2 is  $\approx 2k = \frac{\log n}{\log \nu}$ .

#### DIAMETER

Generating function of  $\{q_k\}_{k=0}^{\infty}$ :  $G_q(z) = \sum_{k=0}^{\infty} q_k z^k$ .

The extinction probability of the branching process with offspring distribution  $\{q_k\}$  is the smallest solution in [0, 1] of the fixed point equation  $\beta = G_q(\beta)$ . We define

$$\beta_* = G'_q(\beta) = \sum_{k=1}^{\infty} k q_k \beta^{k-1}.$$

 $d_{\min}$  is the minimum degree of the graph.

#### Theorem 2. We have

$$\frac{\operatorname{diam}(G(n, (d_i)_1^n))}{\log n} \xrightarrow{p} \frac{1}{\log \nu} - \frac{\mathbf{1}(d_{\min} = 2)}{\log q_1} - 2\frac{\mathbf{1}(d_{\min} = 1)}{\log \beta_*}$$
$$= \frac{1}{\log \nu} + \frac{2 - \mathbf{1}(d_{\min} \ge 2) - \mathbf{1}(d_{\min} \ge 3)}{|\log \beta_*|}$$

Bollobás, de la Vega 1982 (random regular graphs)

#### Fernholz, Ramachandran 2007

#### LOWER DEVIATIONS FOR SUPERCRITICAL GWP

A dichotomy:

 $\begin{array}{ll} \text{Schröder case} \Leftrightarrow & q_0 + q_1 > 0 & \Rightarrow \lim_{n \to \infty} \beta_*^{-n} \mathbb{P}(Z_n = k) = \nu_k, \quad k \geq 1 \\ \text{Böttcher case} \Leftrightarrow & q_0 + q_1 = 0 & \Rightarrow \mu = \min\{j, q_j > 0\} \geq 2 \end{array}$ 

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More in: Fleischmann Wachtel 2005

#### **TYPICAL WEIGHTED DISTANCE**

**Theorem 3.** For *a* and *b* chosen uniformly at random in  $G(n, (d_i)_1^n)$  with  $d_{\min} \ge 2$  and with *i.i.d.* exponential 1 weights on its edges, we have

$$\frac{\operatorname{dist}_w(a,b)}{\log n} \xrightarrow{p} \frac{1}{\nu - 1}$$

Bhamidi, Van der Hofstad, Hooghiemstra 2009

Bhamidi, Van der Hofstad, Hooghiemstra 2010 for Erdős-Rényi random graphs.

#### **TYPICAL WEIGHTED DISTANCE**

**Theorem 4.** For *a* and *b* chosen uniformly at random in  $G(n, (d_i)_1^n)$  with  $d_{\min} \ge 2$  and with *i.i.d.* exponential 1 weights on its edges, we have

$$\frac{\operatorname{dist}_w(a,b)}{\log n} \xrightarrow{p} \frac{1}{\nu - 1}.$$

Bhamidi, Van der Hofstad, Hooghiemstra 2009 Bhamidi, Van der Hofstad, Hooghiemstra 2010 for Erdős-Rényi random graphs. Recall:

$$\frac{\operatorname{dist}(a,b)}{\log n} \xrightarrow{p} \frac{1}{\log \nu} \ge \frac{1}{\nu - 1}.$$

# HEURISTIC: SPLIT TIMES



$$\widehat{S}_k(a) = d_a + \widehat{d}_1 + \dots + \widehat{d}_k - k$$
  
$$T_k(a) = \sum_{i=0}^{k-1} \frac{E_i}{\widehat{S}_i(a)}$$

We have  $\mathbb{E}[\widehat{S}_i(a)] \approx (\nu-1)i$  and

$$\lim_{k \to \infty} \frac{T_k}{\log k} = \frac{1}{\nu - 1}$$

# WEIGHTED DIAMETER

**Theorem 5.** Consider a random graph  $G(n, (d_i)_1^n)$  with *i.i.d.* exponential 1 weights on *its edges, then* 

$$\frac{\operatorname{diam}_w(G(n, (d_i)_1^n))}{\log n} \xrightarrow{p} \frac{1}{\nu - 1} + \frac{2}{d_{\min}} \mathbf{1}_{(d_{\min} \ge 3)} + \frac{\mathbf{1}_{(d_{\min} = 2)}}{1 - q_1} + \frac{2}{1 - \beta_*} \mathbf{1}_{(d_{\min} = 1)}.$$

Ding, Han Kim, Lubetzky, Peres 2010 (random regular graphs) Amini, Draief, L. 2010

# WEIGHTED DIAMETER

**Theorem 6.** Consider a random graph  $G(n, (d_i)_1^n)$  with *i.i.d.* exponential 1 weights on *its edges, then* 

$$\frac{\operatorname{diam}_w(G(n, (d_i)_1^n))}{\log n} \xrightarrow{p} \frac{1}{\nu - 1} + \frac{2}{d_{\min}} \mathbf{1}_{(d_{\min} \ge 3)} + \frac{\mathbf{1}_{(d_{\min} = 2)}}{1 - q_1} + \frac{2}{1 - \beta_*} \mathbf{1}_{(d_{\min} = 1)}.$$

Ding, Han Kim, Lubetzky, Peres 2010 (random regular graphs) Amini, Draief, L. 2010

Recall:

$$\frac{\operatorname{dist}_w(a,b)}{\log n} \xrightarrow{p} \frac{1}{\nu-1}$$

$$\frac{\operatorname{diam}(G(n,(d_i)_1^n))}{\log n} \xrightarrow{p} \frac{1}{\log \nu} - \frac{\mathbf{1}(d_{\min}=2)}{\log q_1} - 2\frac{\mathbf{1}(d_{\min}=1)}{\log \beta_*}.$$

# LARGE DEVIATIONS FOR SPLIT TIMES

$$\widehat{S}_k(a) = d_a + \widehat{d}_1 + \dots + \widehat{d}_k - k$$
  $T_k(a) = \sum_{i=0}^{k-1} \frac{E_i}{\widehat{S}_i(a)}$ 

Laplace transform:

$$\mathbb{E}\left[e^{\theta T_k(a)}\right] = \prod_{i=0}^{k-1} \left(1 + \frac{\theta}{\widehat{S}_i(a) - \theta}\right)$$

Case  $d_{\min} \geq 3$ ,

$$\mathbb{P}(T_{\alpha_n}(a) \ge x \log n) \le \mathbb{E}\left[e^{d_{\min}T_{\alpha_n}(a)}\right] \exp(-xd_{\min}\log n)$$
$$\approx \alpha_n^{\frac{d_{\min}}{d_{\min}-2}} n^{-xd_{\min}}$$
$$\mathbb{P}(T_{\beta_n}(a) - T_{\alpha_n}(a) \ge y \log n) \le \mathbb{E}\left[e^{\theta T_{\beta_n}(a) - T_{\alpha_n}(a)}\right] \exp(-\theta y \log n)$$
$$\approx n^{\theta\left(\frac{1}{2(\nu - 1 - \epsilon)} - y\right)}$$

#### A SIMPLE APPLICATION: ASYNC. BROADCAST IN RRG

Each node has a Poisson clock.

One chunk - PUSH

**Corollary 1.** Let  $G \sim \mathcal{G}(n, r)$  be a random r-regular graph with n vertices. Then w.h.p.

$$ABT(G) = 2\left(\frac{r-1}{r-2}\right)\log n + o(\log n).$$



#### Fountoulakis Panagtotou 2010

# THANK YOU!